### Learning features to compare distributions

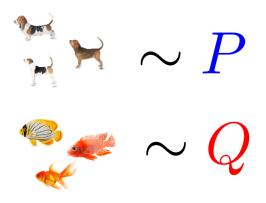
#### **Arthur Gretton**

Gatsby Computational Neuroscience Unit, University College London

Workshop on AI and Neuroscience

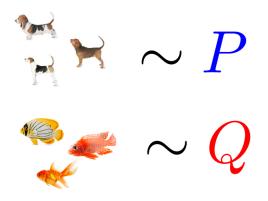
### Goal of this talk

- Have: Two collections of samples X, Y from unknown distributions P and Q.
- Goal: Learn distinguishing features that indicate how P and Q differ.



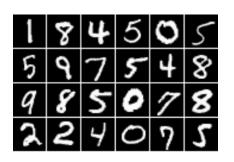
### Goal of this talk

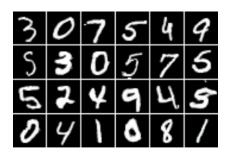
- Have: Two collections of samples X, Y from unknown distributions P and Q.
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### Goal of this talk

- Have: Two collections of samples X, Y from unknown distributions P and Q.
- Goal: Learn distinguishing features that indicate how P and Q differ.

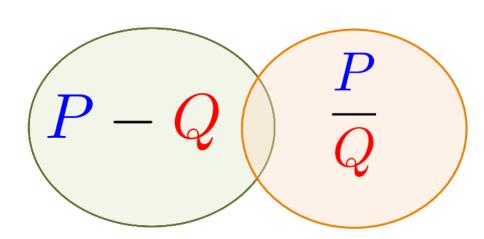


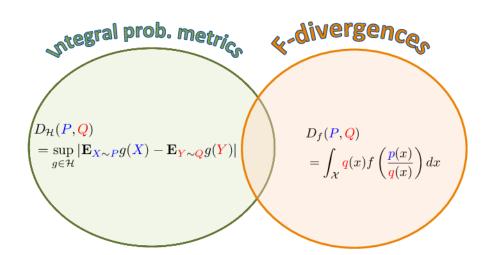


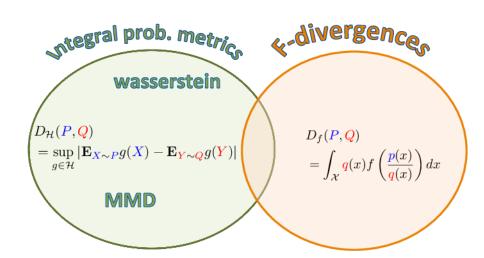
MNIST samples

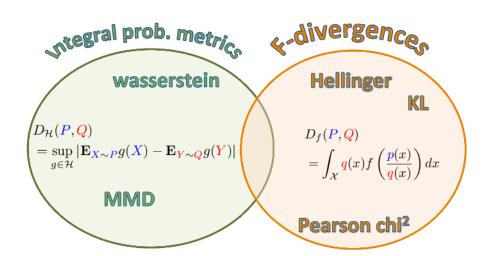
Samples from a GAN

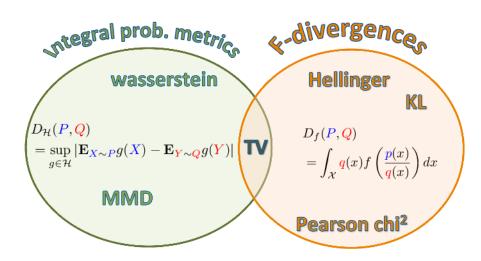
Significant difference in GAN and MNIST? 3/36











Sriperumbudur, Fukumizu, G, Schoelkopf, Lanckriet (2012)

### Overview

#### The maximum mean discrepancy:

- How to compute and interpret the MMD
- "Training" the MMD to maximize test power
- Application to troubleshooting GANs

#### The maximum Stein discrepancy:

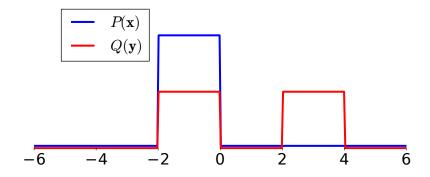
- Divergence between sample and model
- Only need model up to normalizing constant

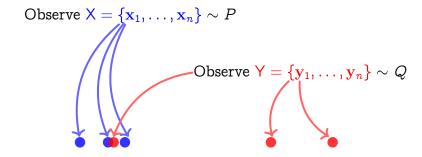
#### The ME test statistic:

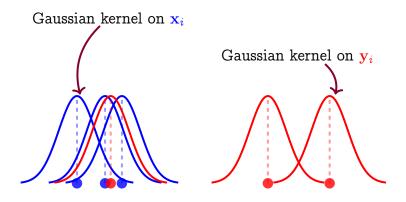
- Informative, linear time features for comparing distributions
- How to learn these features

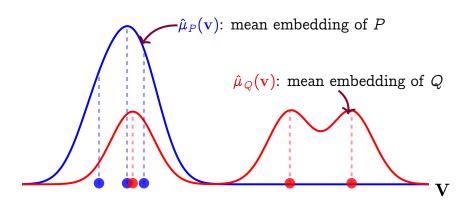
## The maximum mean discrepancy

Are P and Q different?

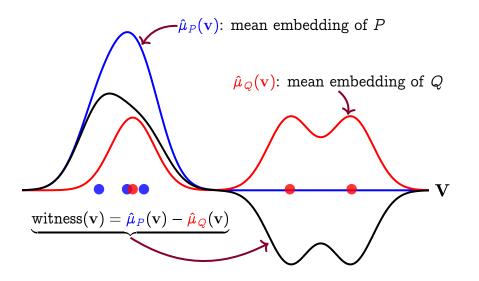


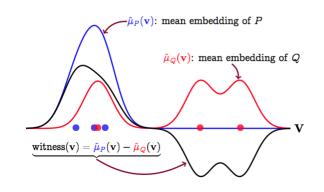






$$\hat{m{\mu}}_P(\mathbf{v}) := rac{1}{m} \sum_{i=1}^m k(\pmb{x}_i, \pmb{v})$$

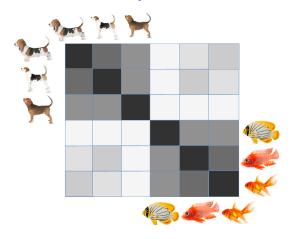




$$egin{aligned} \widehat{MMD}^2 &= \left\| ext{witness}(\mathbf{v}) 
ight\|_{\mathcal{F}}^2 \ &= rac{1}{n(n-1)} \sum_{i 
eq j} k(x_i, x_j) + rac{1}{n(n-1)} \sum_{i 
eq j} k(\mathbf{y}_i, \mathbf{y}_j) \ &- rac{2}{n^2} \sum_{i 
eq j} k(x_i, \mathbf{y}_j) \end{aligned}$$

#### Overview

- Dogs (= P) and fish (= Q) example revisited
- Each entry is one of  $k(dog_i, dog_j)$ ,  $k(dog_i, fish_j)$ , or  $k(fish_i, fish_j)$



#### Overview

The maximum mean discrepancy:

$$\widehat{MMD}^2 = \frac{1}{n(n-1)} \sum_{i \neq j} k(\deg_i, \deg_j) + \frac{1}{n(n-1)} \sum_{i \neq j} k(\operatorname{fish}_i, \operatorname{fish}_j) \\ - \frac{2}{n^2} \sum_{i \neq j} k(\deg_i, \operatorname{fish}_j) \\ k(\operatorname{dog}_i, \operatorname{dog}_j) \quad k(\operatorname{dog}_i, \operatorname{fish}_j)$$

## Asymptotics of MMD

■ The MMD:

$$egin{aligned} \widehat{MMD}^2 = & rac{1}{n(n-1)} \sum_{i 
eq j} k(\pmb{x}_i, \pmb{x}_j) + rac{1}{n(n-1)} \sum_{i 
eq j} k(\pmb{ extbf{y}}_i, \pmb{ extbf{y}}_j) \ & -rac{2}{n^2} \sum_{i,j} k(\pmb{x}_i, \pmb{ extbf{y}}_j) \end{aligned}$$

but how to choose the kernel?

## Asymptotics of MMD

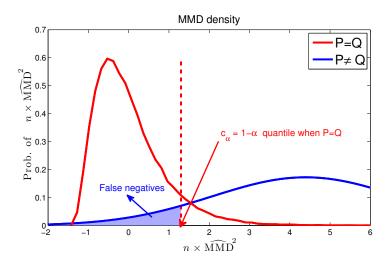
■ The MMD:

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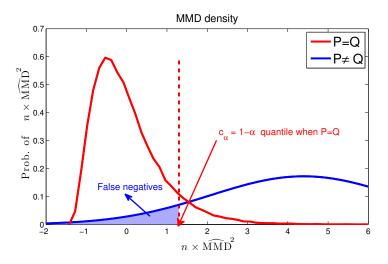
but how to choose the kernel?

- Perspective from statistical hypothesis testing:
  - When P = Q then  $\widehat{MMD}^2$  "close to zero".
  - When  $P \neq Q$  then  $\widehat{MMD}^2$  "far from zero"
- Threshold  $c_{\alpha}$  for  $\widehat{MMD}^2$  gives false positive rate  $\alpha$

### A statistical test

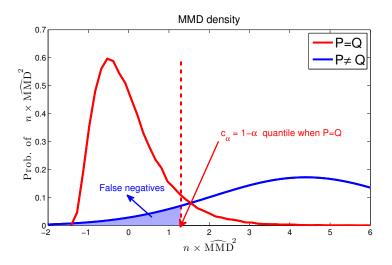


### A statistical test



Best kernel gives lowest false negative rate (=highest power)

### A statistical test



Best kernel gives lowest false negative rate (=highest power)

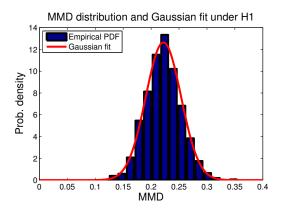
.... but can you train for this?

### Asymptotics of MMD

■ When  $P \neq Q$ , statistic is asymptotically normal,

$$rac{\widehat{ ext{MMD}}^2 - ext{MMD}(P,Q)}{\sqrt{V_n(P,Q)}} \stackrel{\mathcal{D}}{\longrightarrow} \mathcal{N}(0,1),$$

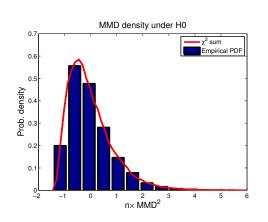
where  $\mathrm{MMD}(P,\,Q)$  is population MMD, and  $\,V_n(P,\,Q)=O\left(n^{-1}
ight)$  .



### Asymptotics of MMD

Where P = Q, statistic has asymptotic distribution

$$n\widehat{ ext{MMD}}^2 \sim \sum_{l=1}^\infty \lambda_l \left[ z_l^2 - 2 
ight]$$



#### where

$$\lambda_i \psi_i(x') = \int_{\mathcal{X}} \underbrace{ ilde{k}(x,x')}_{ ext{centred}} \psi_i(x) dP(x)$$

$$z_l \sim \mathcal{N}(0,2)$$
 i.i.d.

The power of our test (Pr<sub>1</sub> denotes probability under  $P \neq Q$ ):

$$ext{Pr}_1\left(\widehat{n ext{MMD}}^2>\hat{c}_lpha
ight)$$

The power of our test (Pr<sub>1</sub> denotes probability under  $P \neq Q$ ):

$$egin{split} & \Pr_1\left(n\widehat{ ext{MMD}}^2 > \hat{c}_{lpha}
ight) \ & o \Phi\left(rac{ ext{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}} - rac{c_{lpha}}{n\sqrt{V_n(P,Q)}}
ight) \end{split}$$

where

- $\blacksquare$   $\Phi$  is the CDF of the standard normal distribution.
- $\bullet$   $\hat{c}_{\alpha}$  is an estimate of  $c_{\alpha}$  test threshold.

The power of our test (Pr<sub>1</sub> denotes probability under  $P \neq Q$ ):

$$ext{Pr}_1\left(n\widehat{ ext{MMD}}^2>\hat{c}_{lpha}
ight) \ o \Phi\left(\underbrace{rac{ ext{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}}}_{\mathcal{O}(n^{-1/2})}-\underbrace{rac{c_{lpha}}{n\sqrt{V_n(P,Q)}}}_{\mathcal{O}(n^{-3/2})}
ight)$$

Second term asymptotically negligible!

The power of our test (Pr<sub>1</sub> denotes probability under  $P \neq Q$ ):

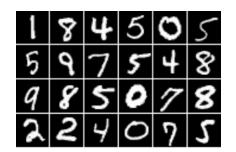
$$egin{split} & ext{Pr}_1 \left( n \widehat{ ext{MMD}}^2 > \hat{c}_{lpha} 
ight) \ & o \Phi \left( rac{ ext{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}} - rac{c_{lpha}}{n \sqrt{V_n(P,Q)}} 
ight) \end{split}$$

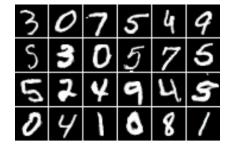
To maximize test power, maximize

$$\frac{\text{MMD}^2(P,Q)}{\sqrt{V_n(P,Q)}}$$

(Sutherland, Tung, Strathmann, De, Ramdas, Smola, G., in review for ICLR 2017) Code: github.com/dougalsutherland/opt-mmd

### Troubleshooting for generative adversarial networks

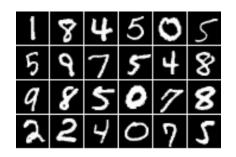




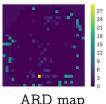
MNIST samples

Samples from a GAN

## Troubleshooting for generative adversarial networks



MNIST samples

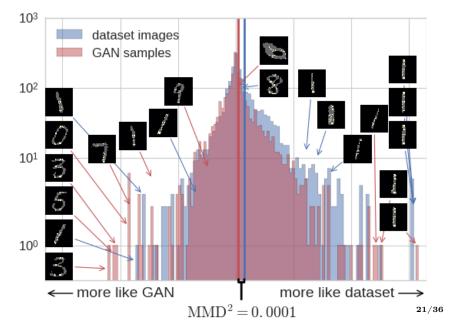


ARD map

Samples from a GAN

- Power for optimzed ARD **kernel**: 1.00 at  $\alpha = 0.01$
- Power for optimized RBF kernel: 0.57 at  $\alpha = 0.01$

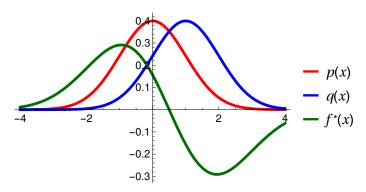
## Benchmarking generative adversarial networks



Testing against a probabilistic model

#### Statistical model criticism

$$MMD(rac{P}{P}, rac{Q}{Q}) = \|f^*\|^2 = \sup_{\|f\|_{\mathcal{F}} \leq 1} [E_Q f - E_{rac{p}{P}} f]$$



 $f^*(x)$  is the witness function

Can we compute MMD with samples from Q and a model P? Problem: usually can't compute  $E_p f$  in closed form.

### Stein idea

To get rid of  $E_{p}f$  in

$$\sup_{\|f\|_{\mathcal{F}} \leq 1} [E_q f - E_{\textcolor{red}{p}} f]$$

we define the Stein operator

$$T_{p}f = \partial_{x}f + f\left(\partial_{x}\log p\right)$$

Then

$$E_{P}T_{P}f=0$$

subject to appropriate boundary conditions. (Oates, Girolami, Chopin, 2016)

#### Stein operator

$$T_{m p}f = \partial_x f + f \partial_x (\log m p)$$

$$MSD(\mathbf{p}, q, \mathcal{F}) = \sup_{\|g\|_{\mathcal{F}} \le 1} E_q T_{\mathbf{p}} g - E_{\mathbf{p}} T_{\mathbf{p}} g$$

#### Stein operator

$$T_{m p}f = \partial_x f + f \partial_x (\log {m p})$$

$$MSD(\mathbf{p}, q, \mathcal{F}) = \sup_{\|g\|_{\mathcal{F}} \le 1} E_q T_{\mathbf{p}} g - E_{\mathbf{p}} T_{\mathbf{p}} g$$

#### Stein operator

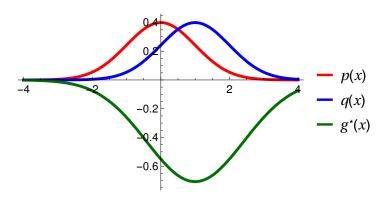
$$T_{m p}f = \partial_x f + f \partial_x (\log {m p})$$

$$MSD({\color{red} p},{\color{gray} q},{\color{gray} {\mathcal F}}) = \sup_{\|{\color{gray} g}\|_{{\color{gray} {\mathcal F}}} \leq 1} E_{{\color{gray} q}} T_{{\color{gray} p}} {\color{gray} g} - {\color{gray} {\it E}}_{{\color{gray} p}} T_{{\color{gray} p}} {\color{gray} g} = \sup_{\|{\color{gray} g}\|_{{\color{gray} {\mathcal F}}} \leq 1} E_{{\color{gray} q}} T_{{\color{gray} p}} {\color{gray} g}$$

#### Stein operator

$$T_{{m p}}f=\partial_x f+f\partial_x(\log {m p})$$

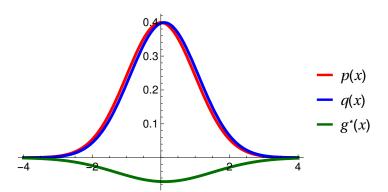
$$MSD({\color{red}p},{\color{gray}q},{\color{gray}\mathcal{F}}) = \sup_{\|{\color{gray}g}\|_{{\color{gray}p}} \leq 1} E_{{\color{gray}q}} T_{{\color{gray}p}} {\color{gray}g} - E_{{\color{gray}p}} T_{{\color{gray}p}} {\color{gray}g} = \sup_{\|{\color{gray}g}\|_{{\color{gray}p}} \leq 1} E_{{\color{gray}q}} T_{{\color{gray}p}} {\color{gray}g}$$



#### Stein operator

$$T_{{m p}}f=\partial_x f+f\partial_x(\log {m p})$$

$$MSD({\color{red}p},{\color{gray}q},{\color{gray}\mathcal{F}}) = \sup_{\|{\color{gray}g}\|_{{\color{gray}p}} \leq 1} E_{{\color{gray}q}} T_{{\color{gray}p}} {\color{gray}g} - E_{{\color{gray}p}} T_{{\color{gray}p}} {\color{gray}g} = \sup_{\|{\color{gray}g}\|_{{\color{gray}p}} \leq 1} E_{{\color{gray}q}} T_{{\color{gray}p}} {\color{gray}g}$$



Closed-form expression for MSD: given  $Z, Z' \sim q$ , then (Chwialkowski, Strathmann, G., 2016) (Liu, Lee, Jordan 2016)

$$\mathrm{MSD}( extbf{\emph{p}},\,q,\mathcal{F}) = E_q h_{ extbf{\emph{p}}}(Z,Z')$$

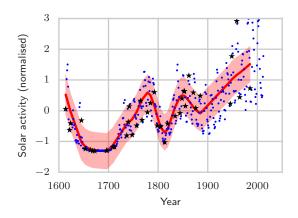
where

$$egin{aligned} h_{m p}(x,y) &:= \partial_x \log m p(x) \partial_x \log m p(y) k(x,y) \ &+ \partial_y \log m p(y) \partial_x k(x,y) \ &+ \partial_x \log m p(x) \partial_y k(x,y) \ &+ \partial_x \partial_y k(x,y) \end{aligned}$$

and k is RKHS kernel for  $\mathcal{F}$ 

Only depends on kernel and  $\partial_x \log p(x)$ . Do not need to normalize p, or sample from it.

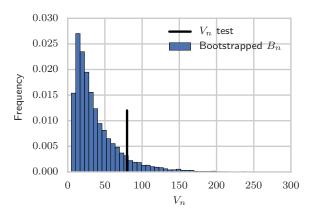
## Statistical model criticism



Test the hypothesis that a Gaussian process model, learned from data \*, is a good fit for the test data (example from Lloyd and Ghahramani, 2015)

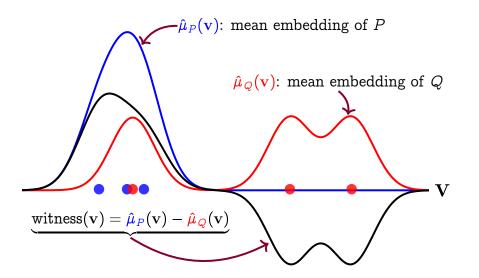
Code: https://github.com/karlnapf/kernel\_goodness\_of\_fit

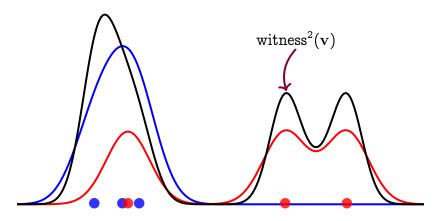
## Statistical model criticism



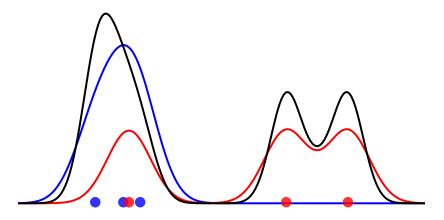
Test the hypothesis that a Gaussian process model, learned from data \*, is a good fit for the test data

# The ME statistic and test

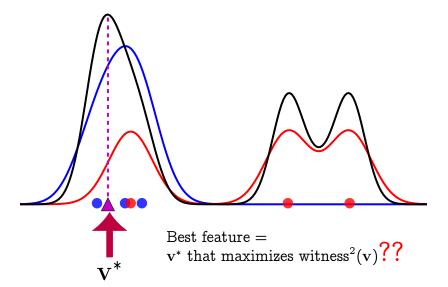


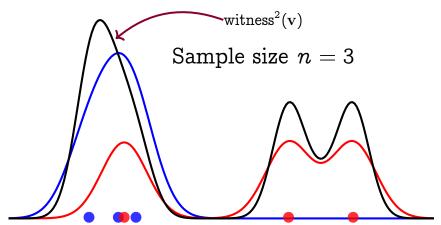


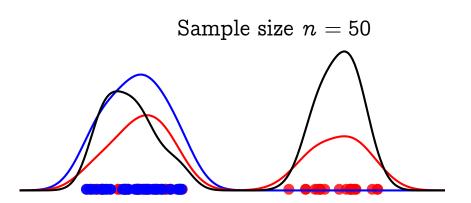
Take square of witness (only worry about amplitude)

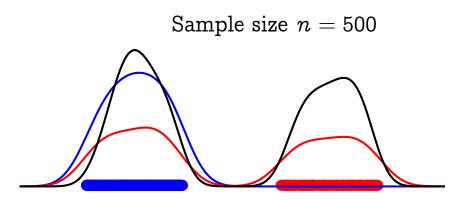


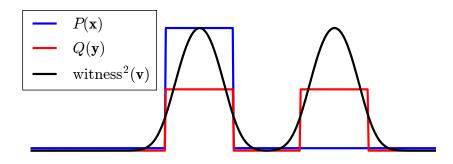
- New test statistic: witness<sup>2</sup> at a single  $\mathbf{v}^*$ ;
- Linear time in number n of samples
- ....but how to choose best feature v\*?



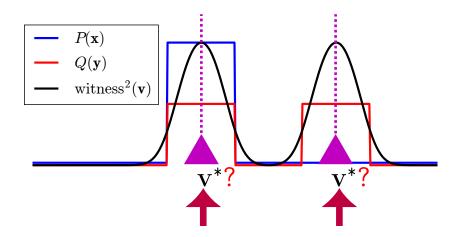








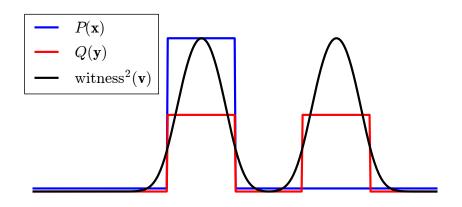
Population witness<sup>2</sup> function



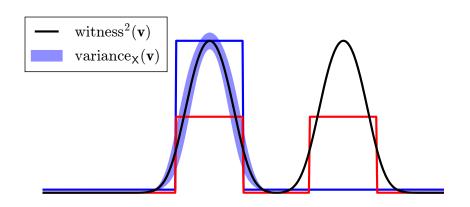
- Variance at  $\mathbf{v} = \text{variance of } X \text{ at } \mathbf{v} + \text{variance of } Y \text{ at } \mathbf{v}$ .
- ME Statistic:  $\hat{\lambda}_n(\mathbf{v}) := n \frac{\text{witness}^2(\mathbf{v})}{\text{variance of } \mathbf{v}}$ .

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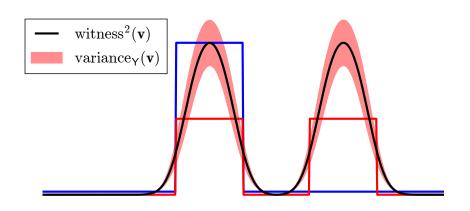
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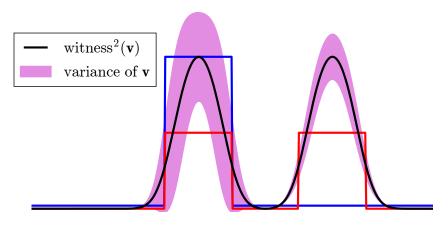
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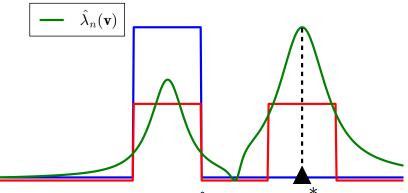
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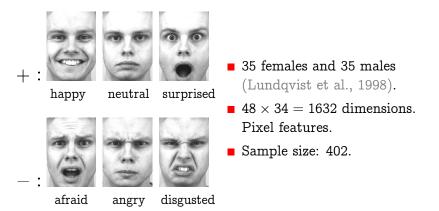
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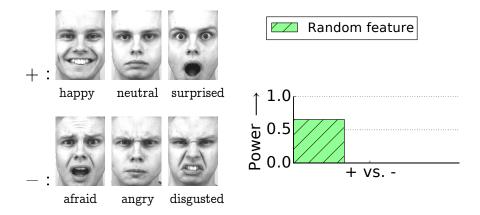
- Variance at  $\mathbf{v} = \text{variance of } X \text{ at } \mathbf{v} + \text{variance of } Y \text{ at } \mathbf{v}$ .
- ME Statistic:  $\hat{\lambda}_n(\mathbf{v}) := n \frac{\text{witness}^2(\mathbf{v})}{\text{variance of } \mathbf{v}}$ .



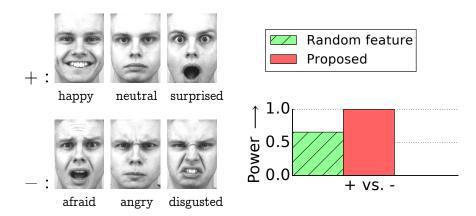
- Best location is  $\mathbf{v}^*$  that maximizes  $\hat{\lambda}_n$ .
- Improve performance using multiple locations  $\{\mathbf{v}_j^*\}_{j=1}^J$



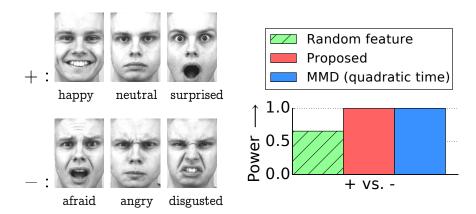
- The proposed test achieves maximum test power in time O(n)
- Informative features: differences at the nose, and smile lines.



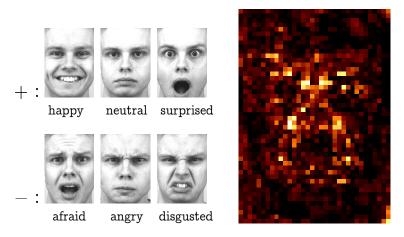
- The proposed test achieves maximum test power in time O(n).
- Informative features: differences at the nose, and smile lines.



- The proposed test achieves maximum test power in time O(n).
- Informative features: differences at the nose, and smile lines.

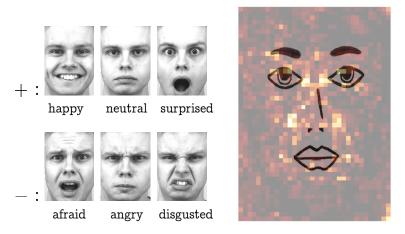


- The proposed test achieves maximum test power in time O(n).
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Learned feature

- The proposed test achieves maximum test power in time O(n).
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Learned feature

- The proposed test achieves maximum test power in time O(n).
- Informative features: differences at the nose, and smile lines.

## Final thoughts

## Witness function approaches:

- MMD test uses pairwise similarities between all samples
- ME test uses similarities to J reference features

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Questions?